

# EFFECT OF MISCLASSIFICATION DUE TO MEASUREMENT ERROR ON THE POWER OF CONTROL CHART FOR PROPORTIONS (ATTRIBUTES)

Ashit B. Chakraborty<sup>1\*</sup>; Anwer Khurshid<sup>2\*\*</sup>

\*Department of Statistics, St. Anthony's College, Shillong, Meghalaya, India

\*\*Department of Mathematical and Physical Sciences, College of Arts and Science, University of Nizwa

## ABSTRACT

The  $p$  chart plays the important role in controlling the proportion of defective items produced. In recent years researchers in various quality control procedures consider the possibility of misclassification errors as an important issue. In the present paper, a method of obtaining the expression of the power of control chart for binomial distribution (proportion of defectives) is being studied by considering approximate expressions for calculating the probabilities of errors of misclassification due to measurement error. Formulae are derived for calculating probabilities of misclassification due to measurement error. The relationship between apparent fraction defective ( $AFD$ ) and true fraction defective ( $TFD$ ) has been used to study the power of control chart.

**KEYWORDS:** Measurement error, misclassification, binomial distribution, power.

**MSC:** 62P30

## RESUMEN

Las p-cartas juegan un rol importante en el control de la proporción de artículos defectuosos producidos. En años recientes investigadores de varios procedimientos de control de la calidad consideran la posibilidad de errores de mala clasificación como un aspecto importante. En el presente trabajo, un método para obtener la expresión de la potencia de la carta de control para la distribución binomial (proporción de defectuosos) ha sido estudiada considerando expresiones aproximadas para calcular las probabilidades de error de mala clasificación debido a errores de medición. Fórmulas son derivadas para calcularlas. La relación entre la aparente fracción de defectuosos ( $AFD$ ) y la verdadera ( $TFD$ ) son usadas para estudiar la potencia de la carta de control.

## 1. INTRODUCTION

Statistical techniques now-a-days have been successfully applied to different production processes in industries to achieve desired quality levels of manufactured products with an optimum production cost. It is widely acknowledged within the industrial process, processes produced are often contaminated with measurement error which can introduce serious bias in the derived results. The nature and magnitude of measurement error and its effect on the actual performance of various control charts can be overwhelming and studied by several researchers. For a recent and brief review see Maravelakis (2012). See also Sankle et al. (2012) and Chakraborty and Khurshid (2013 a, b) and references therein. To employ statistical techniques, inspections are made on the finished products, during the time of production or after the production. In every inspection system, there may be either of two possible types of errors: (i) a good (conforming) item to a specification may be misclassified as defective

---

<sup>1</sup> e-mail: [abc\\_sac@rediffmail.com](mailto:abc_sac@rediffmail.com)

<sup>2</sup> e-mail: [anwer@unizwa.edu.om](mailto:anwer@unizwa.edu.om); [anwer\\_khurshid@yahoo.com](mailto:anwer_khurshid@yahoo.com)

(nonconforming) or (ii) a defective (nonconforming) item may be misclassified as good (conforming). These types of errors are classified as misclassification errors (or inspection errors) and are generally due to chance causes and can be estimated (Sankle and Singh, 2012).

Misclassification is a special case of measurement error. Apparently there is no unified theory that encompasses the key elements of misclassification error which is usually studied separately from measurement error, though there is clearly much overlap. Misclassification errors may significantly alter the performance of (attribute) control charts, as has been investigated by several authors, including Dorris and Foote (1978), Case (1980), Schneider and Tang (1987), Suich (1988) and Johnson et al. (1991). Recently Chen et al. (2011) studied inspection errors in multinomial control charts. More recently Balamurali and Kalyanasundaran (2011) studied the effect of misclassification error on the operating characteristic (OC) curve of analysis of means (ANOM).

The  $p$  chart plays the vital role in controlling the proportion of defective items produced. Singh et al. (2002) illustrated cumulative sum control charts for proportions under inspection error (see also Singh and Sayyed, 2001 for cumulative sum control charts for Poisson variables under inspection error). In the present paper, the power of control chart for binomial distribution (proportion of defectives) is being studied by considering approximate expressions for calculating the probabilities of errors of misclassification due to measurement error. The relationship between apparent fraction defective ( $AFD$ ) and true fraction defective ( $TFD$ ) has been used to study the power of control chart.

## 2. TERMINOLOGY AND FEW ASSUMPTIONS

Summarizing the following notation to be used throughout this paper, will facilitate further development. The misclassification error may be of two types:  $P_1$  (type I error) is the probability of classifying a good item as a defective one and  $P_2$  (type II error) is the probability of classifying a defective item as good one. Further  $AFD$  (apparent fraction defective) is the proportion of defective items if error of misclassification is presented is denoted by  $\pi$  and  $TFD$  (true fraction defective) is the proportion of defective items when there is no error of misclassification and is denoted by  $P$ . It is obvious that  $AFD=TFD$ , if the error of misclassification is zero.

It is assumed that the measurements have been taken only to classify the production items into acceptable and rejectable units with certain specifications that can be expressed in terms of mean and standard deviation of the measurable quality characteristics.

The quality characteristic  $x$  is normally distributed with mean  $\mu$  and standard deviation  $\sigma_p$

$$f(x)dx = \frac{1}{\sigma_p \sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma_p}\right)^2\right] dx.$$

The variable  $v$  is normally distributed with mean  $x$  and standard deviation  $\sigma_e$

$$f(v)dv = \frac{1}{\sigma_e \sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{v-x}{\sigma_e}\right)^2\right] dv.$$

The units beyond  $x = \mu \pm K \sigma_p$  are defective and the units within  $x = \mu \pm K \sigma_p$  are non-defective.

## 3. EVALUATING PROBABILITIES OF MISCLASSIFICATION

Here we have classified the production process, after measurement into one of the two categories. They are either conforming (good) or non-conforming (defective) units. If  $P_1$  is the probability of misclassification of a conforming unit and  $P_2$  is the probability of misclassification of a non-conforming

unit, then owing Singh (1964) with the above mentioned assumptions (Section 2),  $P_1$  and  $P_2$  can be evaluated as

$$P_1 = \int_{-K\sigma_p}^{K\sigma_p} f(x) dx \left[ 1 - \int_{-K\sigma_p}^{K\sigma_p} f(v) dv \right] \quad (1)$$

and

$$P_2 = \int_{K\sigma_p}^{\infty} f(x) dx \int_{-K\sigma_p}^{K\sigma_p} f(v) dv + \int_{-\infty}^{-K\sigma_p} f(x) dx \int_{-K\sigma_p}^{K\sigma_p} f(v) dv. \quad (2)$$

In fact  $P_1$  and  $P_2$  are the inspection risks, which are the type I and type II errors and take the values between 0 and 1.

The approximate expressions for  $P_1$  and  $P_2$  (Singh, 1964) are:

$$P_1 = 2T(h, a) + \{\Phi(k) - \Phi(h)\} \quad (3)$$

and

$$P_2 = 2T(h, a) - \{\Phi(k) - \Phi(h)\} \quad (4)$$

where

$$a = \sigma_e / \sigma_p, \quad h = \frac{K\sigma_p}{\sqrt{\sigma_e^2 + \sigma_p^2}}, \quad \Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x \exp\left[-\frac{1}{2}v^2\right] dv$$

and

$$T(h, a) = \frac{1}{\sqrt{2\pi}} \int_0^a \frac{\exp\left[-\frac{1}{2}h^2(1+x^2)\right]}{1+x^2} dx.$$

Here,  $1.5 \leq K \leq 3$  and  $(\sigma_e / \sigma_p) \leq 0.5$  hold good for finding  $P_1$  and  $P_2$ . Singh (1964) studied measurement error in acceptance sampling plan and calculated  $P_1$  and  $P_2$  based on the graphic representation of the probabilities of misclassification data for different values of  $K$  and  $a = \sigma_e / \sigma_p$ . It has been shown by Lavin (1946) that due to misclassification error, the probability of acceptance of the lot will be obtained by replacing true fraction defective ( $TFD$ )  $P$  by the apparent fraction defective ( $AFD$ )  $\pi$  where

$$\pi = P(1 - P_2) + P_1(1 - P).$$

$\pi$  yields a random variable  $X$  whose binomial distribution has parameter  $\pi$  instead of  $P$ . See also Collins and Case (1976), Johnson et al. (1991) and Mittag and Rinne (1993) for published material based on Lavin equation.

#### 4. POWER OF CONTROL CHART FOR PROPORTIONS DATA

The data can often be represented by a binomial distribution if it consists of a number or proportion of units having a specific attribute. In this section we develop the expression for the power of a control chart for a binomial (p-chart) under misclassification due to measurement error. Recently Khoo (2013) presented power functions for a Shewhart control chart.

Following Kanazuka (1986), the power of detecting the change of process for the control chart is given by

$$P_d = P\{X \geq UCL\} + P\{X \leq LCL\}$$

where  $UCL$  and  $LCL$  are upper and lower control limits respectively.

Thus under misclassification the control limits for a binomial (p-chart) are  $\pi \pm K \sqrt{\frac{\pi(1-\pi)}{n}}$  and

centre line  $CL$  is  $\pi$ . Hence, the power of the control chart under misclassification is

$$P_d = \left[ 1 - \sum_{x_e=0}^{UCL-1} \binom{n}{x_e} \pi^{x_e} (1-\pi)^{n-x_e} \right] + \sum_{x_e=0}^{LCL} \binom{n}{x_e} \pi^{x_e} (1-\pi)^{n-x_e}. \quad (5)$$

The operating characteristic (OC) curve, under misclassification, which illustrates the probability that a sample fraction defective  $x_e/n$ , will fall within control limits as a function of the error process fraction defective  $\pi$  is given by

$$P_e(\pi) = \sum_{x_e=LCL}^{UCL} \binom{n}{x_e} \pi^{x_e} (1-\pi)^{n-x_e}. \quad (6)$$

## 5. EXAMPLE AND ILLUSTRATION

Consider the data for fraction defective (p-chart), where 4 samples, each of size 15 were inspected and number of defectives along with proportions of defectives are obtained as follows:

Sample #	Number of defects ( $d_i$ )	Fraction defectives $P_i = (d_i/n)$
1	1	0.07
2	4	0.27
3	2	0.13
4	5	0.33

Here overall sample proportion of defectives is  $p = \bar{p} = 0.2$  and its standard deviation is

$$\sigma_p = \sqrt{\frac{\bar{p}(1-\bar{p})}{n}} = 0.10. \text{ For our analysis we have kept } p = \bar{p} = 0.2, \text{ the overall sample}$$

proportion of defective fixed and the values of  $n$  being changed in different situations to see the effect of the size of the sample on the power of control chart.

## 6. CALCULATIONS AND CONCLUSIONS

To obtain the power of control chart ( $P_d$ ) and operating characteristic (OC) curve ( $P_e(\pi)$ ) under misclassification error we have to first find  $\pi = P(1 - P_2) + P_1(1 - P)$  based on the approximate expressions for  $P_1$  and  $P_2$  (Equations 3 and 4).

Table 1 gives the values of  $h = \frac{K}{\sqrt{a^2 + 1}}$  for different combinations of  $a = \sigma_e / \sigma_p$  and  $T(h, a)$ .

Here we have used Monte Carlo simulation to find  $T(h, a)$ . True values of fraction defective  $P$  can be obtained from the normal probability table for different values of  $K$ . The values of  $P_1$  and  $P_2$  for different combinations of  $T(h, a)$  and  $\Phi(h)$  for fixed  $K$  have been tabulated in Table 1. It has been observed from Table 1 (A-G) that for fixed  $K$ , the values of  $P_1$  and  $P_2$  show a decreasing trend if the measurement error  $a = \sigma_e / \sigma_p$  decreases. On the other hand, we also observe that for fixed  $a = \sigma_e / \sigma_p$  the values of  $P_1$  is greater than  $P_2$  and when  $h \cong K$  then  $P_1 = P_2$ .

$$\text{Table 1: Values of } T(h, a) = \frac{1}{\sqrt{2\pi}} \int_0^a \frac{\exp\left[-\frac{1}{2}h^2(1+x^2)\right]}{1+x^2} dx, \Phi(h),$$

$$P_1 = 2T(h, a) + \{\Phi(k) - \Phi(h)\} \text{ and } P_2 = 2T(h, a) - \{\Phi(k) - \Phi(h)\}$$

**Table 1-A:** When  $K = 1.5$  and  $\Phi(K) = 0.9332$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.5	1.34	0.07039360	0.9099	0.16408720	0.11748720
0.4	1.39	0.05503907	0.9177	0.12557814	0.09457814
0.3	1.44	0.04001047	0.9251	0.08812094	0.07192094
0.25	1.46	0.03294319	0.9279	0.07118638	0.06058638
0.20	1.47	0.02635462	0.9292	0.05670924	0.04870924
0.15	1.48	0.01970635	0.9306	0.04201270	0.03681270
0.10	1.49	0.01305494	0.9319	0.02740988	0.02480988
0.05	1.50	0.00646451	0.9332	0.01292902	0.01292902

**Table 1-B:** When  $K = 1.75$  and  $\Phi(K) = 0.9599$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.5	1.57	0.04915861	0.9418	0.11641722	0.08021722
0.4	1.63	0.03763664	0.9484	0.08677328	0.06377328
0.3	1.68	0.02722368	0.9535	0.06084736	0.04804736
0.25	1.70	0.02237561	0.9554	0.04925122	0.04025122
0.20	1.72	0.01759671	0.9573	0.03779342	0.03259342
0.15	1.73	0.01315372	0.9582	0.02800744	0.02460744
0.10	1.74	0.008706727	0.9591	0.018213454	0.016613454
0.05	1.75	0.004304809	0.9599	0.008609618	0.008609618

**Table 1-C:** When  $K = 2.0$  and  $\Phi(K) = 0.9772$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.50	1.79	0.03308721	0.9633	0.08007442	0.05227442
0.40	1.86	0.02471443	0.9686	0.05802886	0.04082886
0.30	1.92	0.01745997	0.9726	0.03951994	0.03031994
0.25	1.94	0.01433163	0.9738	0.03206326	0.02526326
0.20	1.96	0.01125022	0.9750	0.02470044	0.02030044
0.15	1.98	0.008244447	0.9761	0.017588894	0.015388894
0.10	1.99	0.00545368	0.9767	0.01140736	0.01040736
0.05	2.00	0.002692772	0.9772	0.005385544	0.005385544

**Table 1-D:** When  $K = 2.25$  and  $\Phi(K) = 0.9878$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.50	2.02	0.020711060	0.9783	0.050922120	0.031922120
0.40	2.09	0.015359960	0.9817	0.036819920	0.024619920
0.30	2.16	0.010555770	0.9846	0.024311540	0.017911540
0.25	2.18	0.008656291	0.9854	0.019712582	0.014912582
0.20	2.20	0.006785243	0.9861	0.015270486	0.011870486
0.15	2.23	0.004852139	0.9871	0.010404278	0.009004278
0.10	2.24	0.003208399	0.9875	0.006716798	0.006116798
0.05	2.25	0.001582424	0.9878	0.003164848	0.003164848

**Table 1-E:** When  $K = 2.50$  and  $\Phi(K) = 0.9938$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.50	2.24	0.012561410	0.9875	0.031422820	0.018822820
0.40	2.33	0.008819303	0.9901	0.021338606	0.013938606
0.30	2.40	0.006015980	0.9918	0.014031960	0.010031960
0.25	2.43	0.004810231	0.9925	0.010920462	0.008320462
0.20	2.45	0.003766157	0.9929	0.008432314	0.006632314
0.15	2.48	0.002681432	0.9934	0.005762864	0.004962864
0.10	2.49	0.001772875	0.9936	0.003745750	0.003345750
0.05	2.50	0.0008734041	0.9938	0.001746808	0.001746808

**Table 1-F:** When  $K = 2.75$  and  $\Phi(K) = 0.9970$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.50	2.47	0.0070560530	0.9932	0.017912106	0.010312106
0.40	2.56	0.0049003870	0.9948	0.012000774	0.007600774
0.30	2.64	0.0032323650	0.9959	0.007564730	0.005364730
0.25	2.67	0.0025776230	0.9962	0.005955246	0.004355246
0.20	2.70	0.0019622790	0.9965	0.004424558	0.003424558
0.15	2.72	0.0014301680	0.9967	0.003160336	0.002560336
0.10	2.74	0.0009200657	0.9969	0.001940131	0.001740131
0.05	2.75	0.0004528698	0.9970	0.000905740	0.000905740

**Table 1-G:** When  $K = 3.00$  and  $\Phi(K) = 0.9987$

$a = \sigma_e / \sigma_p$	$h = \frac{K}{\sqrt{a^2 + 1}}$	$T(h, a)$	$\Phi(h)$	$P_1$	$P_2$
0.50	2.69	0.003860856	0.9964	0.010021712	0.005421712
0.40	2.79	0.002578258	0.9974	0.006456516	0.003856516
0.30	2.88	0.001637422	0.9980	0.003974844	0.002574844
0.25	2.91	0.001302675	0.9982	0.003105350	0.002105350
0.20	2.94	0.000988815	0.9984	0.002277630	0.001677630
0.15	2.97	0.000698636	0.9985	0.001597271	0.001197271

0.10	2.99	0.000448474	0.9986	0.000996947	0.000796947
0.05	3.00	0.000220578	0.9987	0.000441156	0.000441156

**Note:** The function  $T(h, a)$  has been tabulated by Owen (1956) and Smirnov and Bolsev (1962). Interested readers may obtain a simple QBASIC program from the first author.

The relationship between apparent fraction defective ( $AFD$ ) and true fraction defective ( $TFD$ ) is shown in Table 2 and Figure 1.

**Table 2:** Relationship between  $TFD(= P)$  and  $AFD(= \pi)$  when  $K = 2.0$

$P$	$\pi$	$\pi$	$\pi$
	$a = 0.5$	$a = 0.10$	$a = 0.15$
0	0.005385	0.011407	0.017589
0.01	0.015280	0.021119	0.027259
0.02	0.025170	0.030971	0.036929
0.03	0.035116	0.040753	0.046600
0.04	0.044955	0.050535	0.056270
0.05	0.054846	0.060317	0.065940

It is observed that for fixed  $K$  and  $a = \sigma_e / \sigma_p$  as the values of the true fraction defective ( $P$ ) increase, the values of  $\pi$  i.e., apparent (observed) fraction defective also increase and also for fixed  $P$ , as the values of measurement error  $a = \sigma_e / \sigma_p$  increase, there is considerable increase in the values of  $\pi$ .

**Figure 1:** Relationship between apparent fraction defective ( $AFD$ ) and true fraction defective ( $TFD$ )

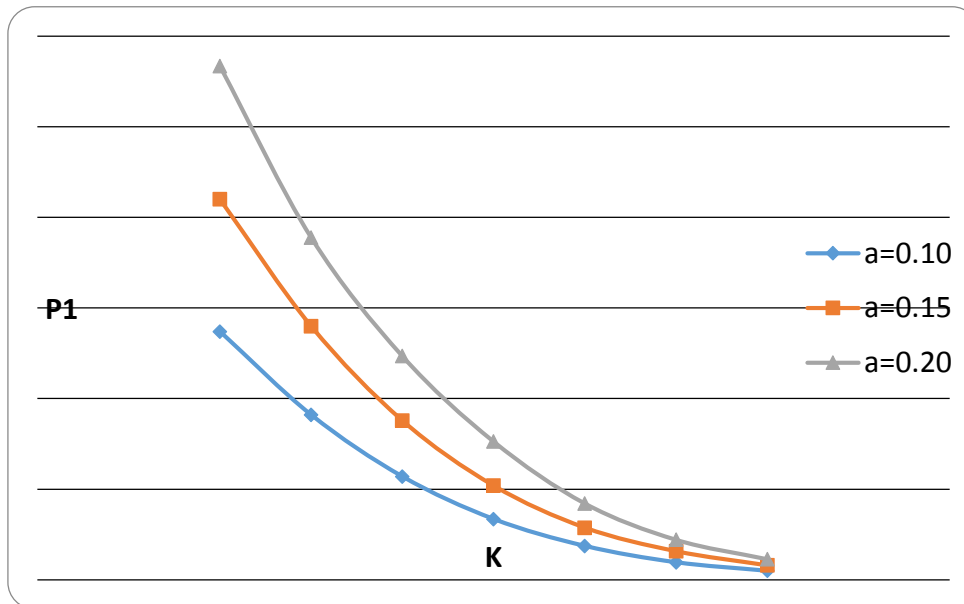
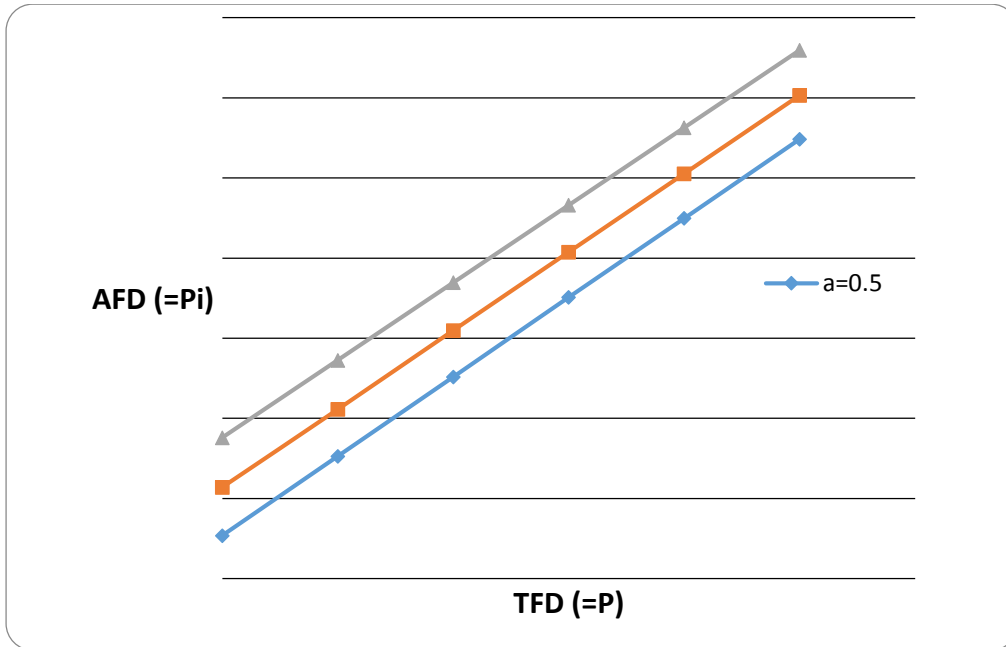


Table 3 and Figure 2 (A and B) depict the effect of  $K$  on probabilities of misclassification of conforming units ( $P_1$ ) and non-conforming units ( $P_2$ ). For fixed  $a = \sigma_e / \sigma_p$ , if we increase the values of  $K$ , there is a decreasing trend for  $P_1$  but for fixed  $K$ , the values of  $P_1$  increase as  $a = \sigma_e / \sigma_p$  is increased. The same trend being observed for Table 3B. shows the graphic representation between  $K$  and probabilities of misclassification. One can also calculate  $P_1$  and  $P_2$  from the graphs (Figure 2) by knowing the standard deviation  $\sigma_e$  of measurement error (which assumes same for all the values of  $K$ ).

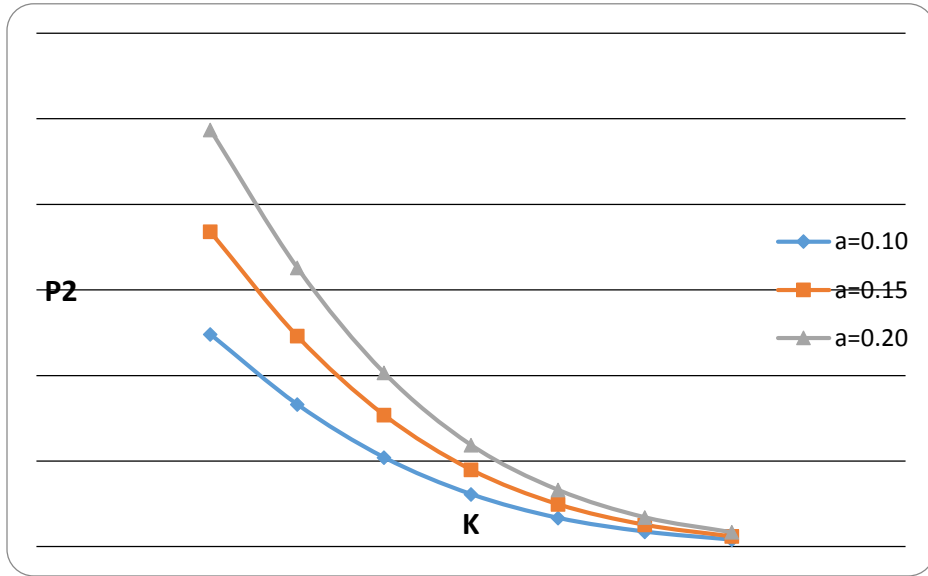
**Table 3:** Probabilities of misclassification of conforming units ( $P_1$ ) and non conforming units ( $P_2$ ) for different values of  $K$  and  $a = \sigma_e / \sigma_p$ .

$K$	$a = 0.10$		$a = 0.15$		$a = 0.20$	
	$P_1$	$P_2$	$P_1$	$P_2$	$P_1$	$P_2$
1.50	0.027409880	0.024809880	0.042012700	0.036812700	0.056709240	0.048709240
1.75	0.018213454	0.016613454	0.028007440	0.024607440	0.037793420	0.032593420
2.00	0.011407360	0.010407360	0.017588894	0.015388894	0.024700440	0.020300440
2.25	0.006716798	0.006116798	0.010404278	0.009004278	0.015270486	0.011870486
2.50	0.003745750	0.003345750	0.005762864	0.004962864	0.008432314	0.006632314
2.75	0.001940131	0.001740131	0.003160336	0.002560336	0.004424558	0.003424558
3.00	0.000996947	0.000796947	0.001597271	0.001197271	0.002277630	0.001677630



**Figure 2 (A):** The effect of  $K$  on probabilities of misclassification of conforming units ( $P_1$ )





**Figure 2 (B):** The effect of  $K$  on probabilities of misclassification of conforming units ( $P_2$ )

Table 4 (A-D) gives us the idea how the values of AFD ( $\pi$ ) values influence the control limits for fraction defective charts. It has been observed from the table that for fixed  $K$ , the values of both  $UCL$  and  $LCL$  increase as there is an increase in the values of  $a = \sigma_e/\sigma_p$ . For fixed  $a = \sigma_e/\sigma_p$ , the difference between  $UCL$  and  $LCL$  increases as we go on increasing  $K$  when the corresponding values of  $\pi$  decrease (which depends on  $P_1$ ,  $P_2$  and  $P$ ). It is observed that the range of  $UCL$  and  $LCL$  is less when the size of the sample is increased.

**Table 4:** Values of  $\pi$  and control limits ( $LCL$  and  $UCL$ ) for different values of  $K$ ,  $P_1$ ,  $P_2$ ,  $a = \sigma_e/\sigma_p$  and  $n$  for fixed  $p = \bar{p} = 0.2$

**Table 4 A:**  $a = \sigma_e/\sigma_p = 0.05$

$K$	$P_1$	$P_2$	$\pi$	$n = 15$		$n = 50$	
				$LCL$	$UCL$	$LCL$	$UCL$
1.50	0.012929020	0.012929020	0.2078	0.0506	0.3649	0.1217	0.2939
1.75	0.008609618	0.008609618	0.2052	0.0227	0.3877	0.1053	0.3051
2.00	0.005385544	0.005385544	0.2032	0	0.4110	0.0894	0.3170
2.25	0.003164848	0.003164848	0.2019	0	0.4351	0.0742	0.3296
2.50	0.001746808	0.001746808	0.2010	0	0.4597	0.0593	0.3427
2.75	0.000905740	0.000905740	0.2006	0	0.4849	0.0449	0.3563
3.00	0.000441156	0.000441156	0.2003	0	0.5102	0.0305	0.3701

**Table 4 B:**  $a = \sigma_e/\sigma_p = 0.15$

$K$	$P_1$	$P_2$	$\pi$	$n = 15$		$n = 50$	
				$LCL$	$UCL$	$LCL$	$UCL$
1.50	0.042012700	0.036812700	0.2262	0.0642	0.3882	0.1375	0.3149
1.75	0.028007440	0.024607440	0.2175	0.0311	0.4039	0.1154	0.3195
2.00	0.017588894	0.015388894	0.2110	0.0003	0.4217	0.0956	0.3264
2.25	0.010404278	0.009004278	0.2065	0	0.4417	0.0777	0.3353
2.50	0.005762864	0.004962864	0.2036	0	0.4635	0.0612	0.3460
2.75	0.003160336	0.002560336	0.2020	0	0.4871	0.0459	0.3581
3.00	0.001597271	0.001197271	0.2010	0	0.5114	0.0310	0.3710

**Table 4 C:**  $a = \sigma_e/\sigma_p = 0.25$

				$n = 15$	$n = 50$

$K$	$P_1$	$P_2$	$\pi$	$LCL$	$UCL$	$LCL$	$UCL$
1.50	0.071186380	0.060586380	0.2448	0.0783	0.4113	0.1536	0.3360
1.75	0.049251220	0.040251220	0.2314	0.0408	0.4220	0.1270	0.3358
2.00	0.032063260	0.025263260	0.2210	0.0067	0.4353	0.1036	0.3384
2.25	0.019712582	0.014912582	0.2128	0	0.4506	0.0826	0.3430
2.50	0.010920462	0.008320462	0.2071	0	0.4687	0.0638	0.3504
2.75	0.005955246	0.004355246	0.2039	0	0.4900	0.0472	0.3606
3.00	0.003105350	0.002105350	0.2021	0	0.5132	0.0317	0.3725

**Table 4 D:**  $a = \sigma_e / \sigma_p = 0.50$

$K$	$P_1$	$P_2$	$\pi$	$n = 15$		$n = 50$	
				$LCL$	$UCL$	$LCL$	$UCL$
1.50	0.164087200	0.117487200	0.3078	0.1290	0.4866	0.2099	0.4057
1.75	0.116417220	0.080217220	0.2771	0.0822	0.4793	0.1663	0.3879
2.00	0.080074420	0.052274420	0.2536	0.0289	0.4783	0.1305	0.3767
2.25	0.050922120	0.031922120	0.2344	0	0.4805	0.0996	0.3692
2.50	0.031422820	0.018822820	0.2214	0	0.4894	0.0746	0.3682
2.75	0.017912106	0.010312106	0.2123	0	0.5027	0.0533	0.3713
3.00	0.010021712	0.005421712	0.2069	0	0.5207	0.0350	0.3788

Table 5 (A-K) shows the different values of power of control chart ( $P_d$ ) for the corresponding values of  $\pi$ . Here we observe how power curve ( $P_d$ ) changes for different values of  $n$ ,  $K$ ,  $a = \sigma_e / \sigma_p$ ,  $UCL$  and  $LCL$ . From the Table 5 (A, B, C) it is observed that values of  $P_d$  go on decreasing as we increase  $K$  ( $K=1.5$  to  $K=3$ ) for fixed  $a = \sigma_e / \sigma_p$  and  $P_1 = P_2$ . Also no change in the values of  $P_d$  being observed if there is marginal increase in the values of  $a = \sigma_e / \sigma_p$  for fixed  $n$  and fixed  $K$ . But if we increase the size of the sample (Table 5 D) for fixed  $K$  and  $P_1 = P_2$  there is a change in the values of  $P_d$ . The values of the power ( $P_d$ ) is less if the size of the sample is larger for fixed  $a = \sigma_e / \sigma_p$ . It is also understood from the Table 5 (E), that the values of the power ( $P_d$ ) is more, if  $n$  increased along with the value of  $a = \sigma_e / \sigma_p$ .

It has been observed from the Table 5 that as we go on increasing the shift of the process parameter, there is an increase in the power of the control chart for fixed  $a = \sigma_e / \sigma_p$ ,  $n$ ,  $K$ ,  $UCL$  and  $LCL$ .

Thus, smaller the deviation, smaller the power of the test. As we increase  $a = \sigma_e / \sigma_p$ , keeping other parameters fixed, it has been observed from the table that  $P_2$  (type II error) value increases for fixed deviation, and  $P_2$  values tend to decrease as there is an increase in deviation. Higher values of  $P_2$  may involve cost. Thus, where it is necessary to have a sample of small size,  $P_d$  should be set at a relatively high level so that the resultant  $P_2$  value does not become a matter of excessive concern. Increase in the sample size  $n$  also shifts the power curve upward. Graphical representation for some values of  $P_d$  for the Table 5 (A, B, C) is shown in Figure 3 for different values of  $K$ .

**Table 5:** Power of control chart for proportions under misclassification (due to measurement error)

<b>Table: 5 A</b>	<b>Table: 5 B</b>	<b>Table: 5 C</b>
$a = \sigma_e / \sigma_p = 0.05$	$a = \sigma_e / \sigma_p = 0.05$	$a = \sigma_e / \sigma_p = 0.05$
$p = \bar{p} = 0.2$	$p = \bar{p} = 0.2$	$p = \bar{p} = 0.2$
$n = 15, K = 1.5$	$n = 15, K = 2$	$n = 15, K = 3$

	$P_1 = P_2 = 0.01292902$ $CL = 3.1, UCL = 5,$ $LCL = 1$	$P_1 = P_2 = 0.005385544$ $CL = 3, UCL = 5, LCL = 0$	$P_1 = P_2 = 0.0004411564$ $CL = 3.005, UCL = 7,$ $LCL = 0$
$\pi$	$P_d$	$P_d$	$P_d$
0.01	0.9904	0.8601	0.8601
0.02	0.9674	0.7386	0.7386
0.04	0.8811	0.5421	0.5421
0.05	0.8296	0.4634	0.4633
0.07	0.7196	0.3370	0.3367
0.09	0.6117	0.2443	0.2432
0.10	0.5617	0.2081	0.2062
0.15	0.3354	0.1042	0.0910
0.20	0.3313	0.0963	0.0533
0.25	0.3937	0.1618	0.0700
0.35	0.6623	0.4373	0.2468
0.45	0.8813	0.7393	0.5479
0.50	0.9413	0.8491	0.6964
0.65	0.9972	0.9876	0.9578
0.75	0.9999	0.9992	0.9958

	<b>Table: 5 D</b> $a = \sigma_e / \sigma_p = 0.05$ $p = \bar{p} = 0.2$ $n = 20, K = 1.5$ $P_1 = P_2 = 0.01292902$ $CL = 4.156,$ $UCL = 7, LCL = 1$	<b>Table: 5 E</b> $a = \sigma_e / \sigma_p = 0.50$ $p = \bar{p} = 0.2$ $n = 20, K = 1.5$ $P_1 = 0.1640872,$ $P_2 = 0.1174872$ $CL = 6.16,$ $UCL = 9, LCL = 3$	<b>Table: 5 F</b> $a = \sigma_e / \sigma_p = 0.50$ $p = \bar{p} = 0.2$ $n = 20, K = 3$ $P_1 = 0.010021712,$ $P_2 = 0.005421712$ $CL = 4, UCL = 9,$ $LCL = 0$
$\pi$	$P_d$	$P_d$	$P_d$
0.01	0.9831	1.0000	0.8179
0.02	0.9401	0.9994	0.6676
0.04	0.8103	0.9926	0.4420
0.05	0.7358	0.9841	0.3585
0.10	0.3941	0.8671	0.1217
0.20	0.1559	0.4214	0.0215
0.25	0.2385	0.2661	0.0441
0.35	0.5855	0.2820	0.2378
0.40	0.7505	0.4204	0.4044
0.45	0.8702	0.5906	0.5857
0.50	0.9432	0.7496	0.7483
0.65	0.9887	0.9804	0.9804

	<b>Table: 5 G</b> $a = \sigma_e / \sigma_p = 0.15$ $p = \bar{p} = 0.2$ $n = 50, K = 1.5$ $P_1 = 0.0420127, P_2 = 0.0368127$ $CL = 11.31, UCL = 19, LCL = 3$	<b>Table: 5 H</b> $a = \sigma_e / \sigma_p = 0.15$ $p = \bar{p} = 0.2$ $n = 50, K = 3$ $P_1 = 0.0015972714,$ $P_2 = 0.0011972714$ $CL = 10, UCL = 25, LCL = 0$
--	--	--

$\pi$	$P_d$	$P_d$
0.01	0.9984	0.6050
0.02	0.9822	0.3642
0.05	0.7604	0.0769
0.07	0.5327	0.0266
0.10	0.2503	0.0052
0.20	0.0082	0
0.25	0.0293	0.0001
0.35	0.3784	0.0207
0.40	0.6644	0.0978
0.50	0.9675	0.5561
0.60	0.99948	0.942656
0.65	0.99997	0.989956
0.75	0.9999	0.999962

	<b>Table: 5 I</b> (when $AFD = TFD$ ) $\pi = p = \bar{p} = 0.2$ $n = 15, K = 1.5$ $CL = 3, UCL = 5,$ $LCL = 1$	<b>Table: 5 J</b> (when $AFD = TFD$ ) $\pi = p = \bar{p} = 0.2$ $n = 20, K = 1.5$ $CL = 3,$ $UCL = 6, LCL = 1$	<b>Table: 5 K</b> (when $AFD = TFD$ ) $\pi = p = \bar{p} = 0.2$ $n = 50, K = 3$ $CL = 3, UCL = 18,$ $LCL = 2$
$\pi = p$	$P_d$	$P_d$	$P_d$
0.01	0.9904	0.9904	0.9862
0.02	0.9647	0.9647	0.9216
0.05	0.8296	0.7361	0.5404
0.10	0.5617	0.4030	0.1117
0.20	0.3313	0.2650	0.0004
0.25	0.3937	0.4071	0.0552
0.35	0.6623	0.7567	0.4940
0.40	0.7879	0.8749	0.7631
0.45	0.8813	0.9948	0.9235
0.50	0.9413	0.9793	0.9836
0.65	0.9973	0.9997	0.9999

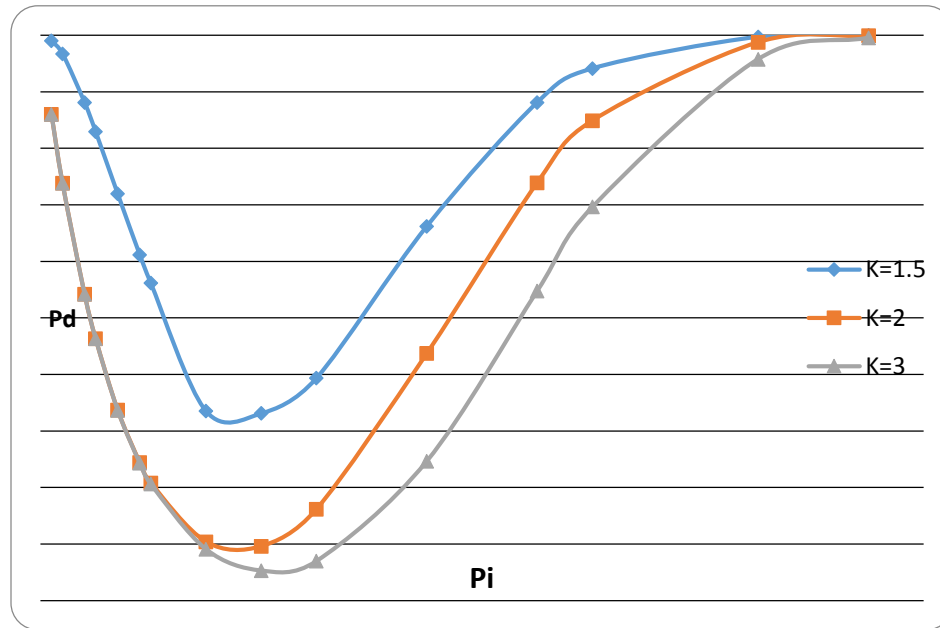


Figure 3: Relationship between  $\pi$  and  $P_d$  for the Table 5 (A, B, C) for different values of

RECEIVED: FEBRUARY 2015  
 REVISED: JANUARY 2016

#### REFERENCES

- [1] BALAMURALI, S. and KALYANASUNDARAM, M. (2011). An investigation of the effects of misclassification errors on the analysis of means. **Tamsui Oxford Journal of Information and Mathematical Sciences**, 27, 117-136.
- [2] CASE, K. E. (1980). The  $p$  control chart under inspection error. **Journal of Quality Technology**, 12, 1-9.
- [3] CHAKRABORTY, A. B. and KHURSHID, A. (2013 a). Measurement error effect on the power of control chart for zero-truncated Poisson distribution. **International Journal for Quality Research**, 7, 411-419.
- [4] CHAKRABORTY, A. B. and KHURSHID, A. (2013 b). Measurement error effect on the power of control chart for the ratio of two Poisson distributions. **Economic Quality Control**, 28, 15-21.
- [5] CHEN, L-H., CHANG, F. M. and CHEN, Y-L. (2011). The application of multinomial control charts for inspection error. **International Journal of Industrial Engineering**, 18, 244-253.
- [6] COLLINS, R. D. and CASE, K. E. (1976). The distribution of observed defectives in attribute acceptance sampling plans under inspection errors. **AIIE Transactions**, 8, 375-378.
- [7] DORRIS, A. L. and FOOTE, B. L. (1978). Inspection errors and statistical quality control: A survey. **AIIE Transactions**, 10, 184-192.
- [8] JOHNSON, N. L., KEMP, A. W. and KOTZ, S. (2005). **Univariate Discrete Distributions**, Third Edition, John Wiley, New York.
- [9] JOHNSON, N. L., KOTZ, S. and WU, X. (1991). **Inspection Error for Attributes in Quality Control**. Chapman and Hall, New York.
- [10] KANAZUKA, T. (1986). The effects of measurement error on the power of  $\bar{X} - R$  charts. **Journal of Quality Technology**, 18, 91-95.
- [11] KHOO, M. B. C. (2013). Power functions of the Shewhart control chart. **Journal of Physics: Conference Series**, 423, 1-5.
- [12] LAVIN, M. (1946). Inspection and efficiency and sampling inspection plans. **Journal of American Statistical Association**, 41, 432-438.

- [13] MARAVELAKIS, P. E. (2012). Measurement error effect on the CUSUM control chart. **Journal of Applied Statistics**, 39, 323-336.
- [14] MITTAG, H. -J. and RINNE, D. (1993). **Statistical Methods of Quality Assurance**, Second Edition, Chapman and Hall/CRC, New York.
- [15] OWEN, D. B. (1956). Tables for computing bivariate normal probabilities. **Annals of Mathematical Statistics**, 27, 1075-1090.
- [16] SANKLE, R. and SINGH, J. R. (2012). Single sampling plans for variables indexed by AQL and AOQL with measurement error. **Journal of Modern Applied Statistical Methods**, 11, 396-406.
- [17] SANKLE, R., SINGH, J. R. and MANGAL, I. K. (2012). Cumulative sum control charts for truncated normal distribution under measurement error. **Statistics in Transition (New Series)**, 13, 95-106.
- [18] SCHNEIDER, H. and TANG, K. (1987). The effects of inspection error on complete inspection plan. **AIIE Transactions**, 19, 421-428.
- [19] SINGH, H. R. (1964). Measurement error in acceptance sampling for attributes. **Indian Society of Quality Control Bulletin**, x, 29-36.
- [20] SINGH, J. R. and SAYYED, M. (2001). Cumulative sum control chart for Poisson variables under inspection error. **Varahmihir Journal of Mathematical Sciences**, 1, 203-209.
- [21] SINGH, J. R., SAYYED, M. and SONI, D. (2002). Cumulative sum control chart for proportion under inspection error. **Ultra Science**, 14, 252-261.
- [22] SMIRNOV, N. V. and BOLSEV, L. N. (1962). **Tables for evaluating function of a two dimensional normal distribution**. Izdat.Akad.Nauk.SSSR, Moscow.
- [23] SUICH, R. (1988). The  $c$  control chart under inspection error. **Journal of Quality Technology**, 20, 263-266.