

Correction, Interrogation 2. Sujet A

Exo 1.

$$1. L(\lambda) = \prod_{i=1}^n P(X=X_i) = \left(1 - \frac{1}{\lambda}\right)^{\sum (X_i - 1)} \left(\frac{1}{\lambda}\right)^n = \left(\frac{\lambda-1}{\lambda}\right)^{\sum X_i - n} \lambda^{-n}$$

$$H(\lambda) = \ln L(\lambda) = (\sum X_i - n) (\ln(\lambda-1) - \ln \lambda) - n \ln \lambda$$

$$\frac{dH(\lambda)}{d\lambda} = (\sum X_i - n) \left(\frac{1}{\lambda-1} - \frac{1}{\lambda}\right) - \frac{n}{\lambda} = 0 \Rightarrow \hat{\lambda}_{ML} = \frac{1}{n} \sum X_i = \bar{X}_n$$

$$\frac{d^2 H(\lambda)}{d\lambda^2} = (\sum X_i - n) (-(\lambda-1)^{-2} + \lambda^{-2}) + n\lambda^{-2} = \frac{n\lambda^2 - 2\lambda \sum X_i + \sum X_i^2}{\lambda^2 (\lambda-1)^2}$$

Si $\lambda = \frac{1}{n} \sum X_i$, alors $n\lambda^2 - 2\lambda \sum X_i + \sum X_i^2 = \sum X_i^2 - \frac{2(\sum X_i)^2}{n} < 0$

puisque $X_i \geq 1 \Rightarrow \sum X_i \geq n \Rightarrow \frac{\sum X_i}{n} \geq 1$.

Donc $\frac{d^2 H(\lambda)}{d\lambda^2} < 0$ pour $\lambda = \bar{X}_n$.

2. D'après le théorème central limite, si $E(X_i^2) < \infty$, alors

$$\frac{\bar{X}_n - EX_i}{\sqrt{\frac{\text{Var}(X_i)}{n}}} \Rightarrow \mathcal{N}(0, 1).$$

C'est-à-dire $\frac{\hat{\lambda}_{ML} - \lambda}{\sqrt{\frac{\lambda(\lambda-1)}{n}}} \Rightarrow \mathcal{N}(0, 1).$

On en déduit, ^{que} pour n grand $\hat{\lambda}_{ML}$ suit la loi gaussienne d'espérance λ et de variance $\frac{\lambda(\lambda-1)}{n}$, i.e. $\mathcal{N}\left(\lambda, \frac{\lambda(\lambda-1)}{n}\right)$.

#. Notons $Y \sim \mathcal{N}(0,1)$ et $P(|Y| \leq q_\alpha) = \alpha$.

On a
$$P\left(\frac{|\hat{\lambda}_{uv} - \lambda|}{\sqrt{\frac{\lambda(\alpha-1)}{n}}} \leq q_\alpha\right) = \alpha.$$

① On en déduit
$$P\left(\hat{\lambda}_{uv} - q_\alpha \sqrt{\frac{\lambda(\alpha-1)}{n}} \leq \lambda \leq \hat{\lambda}_{uv} + q_\alpha \sqrt{\frac{\lambda(\alpha-1)}{n}}\right) = \alpha,$$

donc
$$IC_\alpha = \left[\bar{X}_n - q_\alpha \sqrt{\frac{\lambda(\alpha-1)}{n}}, \bar{X}_n + q_\alpha \sqrt{\frac{\lambda(\alpha-1)}{n}} \right].$$

②
$$-q_\alpha \leq \frac{\sqrt{n}(\hat{\lambda}-1)}{\lambda} \leq q_\alpha \Leftrightarrow \frac{\hat{\lambda}}{1 + \frac{q_\alpha}{\sqrt{n}}} \leq \lambda \leq \frac{\hat{\lambda}}{1 - \frac{q_\alpha}{\sqrt{n}}}$$

#.3 D'après l'énoncé, $P(|Y| \leq 1.96) = 0.95$, i.e. $q_{2.5\%} = 1.96$

① On a donc
$$IC_{95\%} = \left[\bar{X}_n - 1.96 \frac{\hat{\lambda}}{\sqrt{n}}, \bar{X}_n + 1.96 \frac{\hat{\lambda}}{\sqrt{n}} \right]$$

②
$$-1.96 \leq \frac{\sqrt{n}(\bar{X}-\lambda)}{\lambda} \leq 1.96 \Leftrightarrow \frac{\bar{X}}{1 + \frac{1.96}{\sqrt{n}}} \leq \lambda \leq \frac{\bar{X}}{1 - \frac{1.96}{\sqrt{n}}}$$

Exo 2.

1. Oui, parce que $f_\theta(x) \geq 0$ et $\int f_\theta(x) dx = \int_0^\theta \frac{2x}{\theta^2} dx = \left[\frac{x^2}{\theta^2} \right]_0^\theta = 1.$

2.
$$F_x(x) = \int_0^x \frac{2t}{\theta^2} dt = \left[\frac{t^2}{\theta^2} \right]_0^x = \frac{x^2}{\theta^2}, \quad x \in [0, \theta].$$

$$F_{\max}(x) = P(\max(X_1, X_2) \leq x) = P(X_1 \leq x, X_2 \leq x) = (P(X_1 \leq x))^2 = \frac{x^4}{\theta^4}, \quad x \in [0, \theta].$$

$$g_\theta(x) = F'_{\max}(x) = \frac{4x^3}{\theta^4}, \quad x \in [0, \theta].$$

$$3. \mathbb{E} X_1 = \int_0^\theta \frac{2x^2}{\theta^2} dx = \left[\frac{2x^3}{3\theta^2} \right]_0^\theta = \frac{2}{3}\theta$$

$$\mathbb{E} \max(X_1, X_2) = \int_0^\theta \frac{4x^4}{\theta^4} dx = \left[\frac{4x^5}{5\theta^4} \right]_0^\theta = \frac{4}{5}\theta$$

$$\mathbb{E} \hat{\theta}_1 = \frac{3}{4} (\mathbb{E} X_1 + \mathbb{E} X_2) = \frac{3}{2} \mathbb{E} X_1 = \theta$$

$$\mathbb{E} \hat{\theta}_2 = \frac{5}{4} \mathbb{E} \max(X_1, X_2) = \theta$$

$$4. \mathbb{E} X_1^2 = \int_0^\theta \frac{2x^3}{\theta^2} dx = \left[\frac{x^4}{2\theta^2} \right]_0^\theta = \frac{1}{2}\theta^2$$

$$\mathbb{E} (\max(X_1, X_2))^2 = \int_0^\theta \frac{4x^5}{\theta^4} dx = \left[\frac{2x^6}{3\theta^4} \right]_0^\theta = \frac{2}{3}\theta^2$$

$$\text{Var}(X_1) = \frac{1}{2}\theta^2 - \left(\frac{2}{3}\theta\right)^2 = \frac{1}{18}\theta^2$$

$$\text{Var}(\max(X_1, X_2)) = \frac{2}{3}\theta^2 - \left(\frac{4}{3}\theta\right)^2 = \left(\frac{2}{3} - \frac{16}{25}\right)\theta^2$$

$$\text{RQM}(\hat{\theta}_1) = \text{Var}(\hat{\theta}_1) = \frac{9}{16} (\text{Var}(X_1) + \text{Var}(X_2)) = \frac{9}{8} \text{Var}(X_1) = \frac{1}{16}\theta^2$$

$$\text{RQM}(\hat{\theta}_2) = \text{Var}(\hat{\theta}_2) = \frac{25}{16} \text{Var}(\max(X_1, X_2)) = \frac{25}{16} \left(\frac{2}{3} - \frac{16}{25}\right)\theta^2 = \frac{1}{24}\theta^2$$

5. Puisque $\text{RQM}(\hat{\theta}_1) > \text{RQM}(\hat{\theta}_2)$, $\hat{\theta}_2$ est meilleur.

Correction, Interrogation 2. Sujet B

Exo 1.

$$1. L(\lambda) = \prod_{i=1}^n f_{\lambda}(x_i) = \left(\frac{1}{\lambda}\right)^n e^{-\frac{1}{\lambda} \sum x_i} = \lambda^{-n} e^{-\frac{1}{\lambda} \sum x_i}$$

$$H(\lambda) = -n \ln \lambda - \frac{1}{\lambda} \sum x_i$$

$$\frac{dH(\lambda)}{d\lambda} = -\frac{n}{\lambda} + \lambda^{-2} \sum x_i = 0 \Rightarrow \hat{\lambda}_{ML} = \frac{1}{n} \sum x_i = \bar{X}_n$$

$$\frac{d^2 H(\lambda)}{d\lambda^2} = \lambda^{-2} n - 2\lambda^{-3} \sum x_i$$

$$\text{Si } \lambda = \frac{1}{n} \sum x_i, \text{ alors } \lambda^{-2} n - 2\lambda^{-3} \sum x_i = -\frac{n^3}{(\sum x_i)^2} < 0.$$

$$\text{Donc } \frac{d^2 H(\lambda)}{d\lambda^2} < 0 \text{ pour } \lambda = \bar{X}_n.$$

2. D'après le théorème central limite, si $E(X_i^2) < \infty$ alors

$$\frac{\bar{X}_n - EX_1}{\sqrt{\frac{\text{Var}(X_1)}{n}}} \Rightarrow \mathcal{N}(0, 1)$$

$$\text{c'est-à-dire } \frac{\hat{\lambda}_{ML} - \lambda}{\sqrt{\frac{\lambda^2}{n}}} \Rightarrow \mathcal{N}(0, 1).$$

On en déduit que pour n grand $\hat{\lambda}_{ML}$ suit la loi gaussienne

d'espérance λ et de variance $\frac{\lambda^2}{n}$, i.e. $\mathcal{N}(\lambda, \frac{\lambda^2}{n})$.

Notons $Y \sim \mathcal{N}(0,1)$ et $P(|Y| \leq q_\alpha) = \alpha$.

$$\text{On a } P\left(\frac{|\hat{\lambda}_{MV} - \lambda|}{\frac{\lambda}{\sqrt{n}}} \leq q_\alpha\right) = \alpha.$$

$$\textcircled{1} \text{ On en déduit } P\left(\hat{\lambda}_{MV} - q_\alpha \frac{\lambda}{\sqrt{n}} \leq \lambda \leq \hat{\lambda}_{MV} + q_\alpha \frac{\lambda}{\sqrt{n}}\right) = \alpha,$$

$$\text{donc } IC_\alpha = \left[\bar{X}_n - \frac{\lambda q_\alpha}{\sqrt{n}}, \bar{X}_n + \frac{\lambda q_\alpha}{\sqrt{n}}\right].$$

$$\textcircled{2} \frac{\hat{\lambda}}{1 + \frac{q_\alpha}{\sqrt{n}}} \leq \lambda \leq \frac{\hat{\lambda}}{1 - \frac{q_\alpha}{\sqrt{n}}}$$

4.3 D'après l'énoncé, $P(|Y| \leq 1.64) = 0.9$, i.e. $q_{90\%} = 1.64$

$$\textcircled{1} \text{ On a donc } IC_{90\%} = \left[\bar{X}_n - \frac{\hat{\lambda} 1.64}{\sqrt{n}}, \bar{X}_n + \frac{\hat{\lambda} 1.64}{\sqrt{n}}\right]$$

$$= [1000 - 164, 1000 + 164].$$

$$\textcircled{2} \frac{\bar{x}}{1 + \frac{1.64}{\sqrt{n}}} \leq \lambda \leq \frac{\bar{x}}{1 - \frac{1.64}{\sqrt{n}}}$$

Exo 2.

$$\begin{aligned} 1. \text{ Oui, parce que } f_0(x) \geq 0 \text{ et } \int_0^\theta f_0(x) dx &= \int_0^\theta \frac{3x^2}{\theta^3} dx \\ &= \left[\frac{x^3}{\theta^3}\right]_0^\theta = 1. \end{aligned}$$

$$2. F_x(x) = \int_0^x \frac{3t^2}{\theta^3} dt = \left[\frac{t^3}{\theta^3}\right]_0^x = \frac{x^3}{\theta^3}, \quad x \in [0, \theta].$$

$$\begin{aligned} F_{\max}(x) &= P(\max(X_1, X_2) \leq x) = P(X_1 \leq x, X_2 \leq x) = (P(X_1 \leq x))^2 \\ &= \frac{x^6}{\theta^6}, \quad x \in [0, \theta]. \end{aligned}$$

$$g_0(x) = F'_{\max}(x) = \frac{6x^5}{\theta^6}, \quad x \in [0, \theta].$$

$$3. \mathbb{E} X_1 = \int_0^\theta \frac{3x^3}{\theta^3} dx = \left[\frac{3x^4}{4\theta^3} \right]_0^\theta = \frac{3}{4}\theta$$

$$\mathbb{E} \max(X_1, X_2) = \int_0^\theta \frac{6x^6}{\theta^6} dx = \left[\frac{6x^7}{7\theta^6} \right]_0^\theta = \frac{6}{7}\theta$$

$$\mathbb{E} \hat{\theta}_1 = \frac{2}{3} (\mathbb{E} X_1 + \mathbb{E} X_2) = \frac{4}{3} \mathbb{E} X_1 = \theta$$

$$\mathbb{E} \hat{\theta}_2 = \frac{7}{6} \mathbb{E} \max(X_1, X_2) = \theta$$

$$4. \mathbb{E} X_1^2 = \int_0^\theta \frac{3x^4}{\theta^3} dx = \left[\frac{3x^5}{5\theta^3} \right]_0^\theta = \frac{3}{5}\theta^2$$

$$\mathbb{E} (\max(X_1, X_2))^2 = \int_0^\theta \frac{6x^7}{\theta^6} dx = \left[\frac{6x^8}{8\theta^6} \right]_0^\theta = \frac{3}{4}\theta^2$$

$$\text{Var}(X_1) = \frac{3}{5}\theta^2 - \left(\frac{3}{4}\theta\right)^2 = \frac{3}{80}\theta^2$$

$$\text{Var}(\max(X_1, X_2)) = \frac{3}{4}\theta^2 - \left(\frac{6}{7}\theta\right)^2 = \left(\frac{3}{4} - \frac{36}{49}\right)\theta^2$$

$$\text{RQM}(\hat{\theta}_1) = \text{Var}(\hat{\theta}_1) = \frac{4}{9} (\text{Var}(X_1) + \text{Var}(X_2)) = \frac{8}{9} \text{Var}(X_1) = \frac{1}{30}\theta^2$$

$$\text{RQM}(\hat{\theta}_2) = \text{Var}(\hat{\theta}_2) = \frac{49}{36} \text{Var}(\max(X_1, X_2)) = \frac{49}{36} \left(\frac{3}{4} - \frac{36}{49}\right)\theta^2 = \frac{1}{48}\theta^2$$

5. Puisque $\text{RQM}(\hat{\theta}_1) > \text{RQM}(\hat{\theta}_2)$, $\hat{\theta}_2$ est meilleur.