

Limit theory for quadratic forms of linear processes

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Problem: what simple verifiable condition for the CLT for general quadratic form of linear variables

Sums Assume that (X_t) is a stationary linear process

$$X_t = \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j},$$

where (ε_t) is a sequence of i.i.d. $(0,1)$ variables

$$E\varepsilon_t = 0, \quad E\varepsilon_t^2 = 1$$

$$\psi_j \text{ real} \quad \sum_{j=0}^{\infty} \psi_j^2 < \infty.$$

Dependence: Variables X_t can be weakly and strongly dependent

Asymptotic theory for the sums: If

$$S_n = \sum_{j=1}^n X_j, \quad \text{and} \quad \text{Var}(S_n) \rightarrow \infty$$

then (Well known (Ibragimov, Linnik (1964))):

$$\frac{S_n - E[S_n]}{\sqrt{\text{Var}(S_n)}} \rightarrow N(0, 1).$$

Note: no additional conditions needed,

Quadratic forms in i.i.d. variables

Question: can we have something like that for quadratic forms?

Answer: Yes, for i.i.d variables $X_t = \varepsilon_t$. Set

$$T_n = \sum_{t,s=1}^n a_n(t,s) \varepsilon_t \varepsilon_s$$

$A_n = (a_n(t,s))$ real symmetric matrix

Question: Does

$$\text{Var}(T_n) \rightarrow \infty$$

implies CLT for T_n ?

Answer: Almost: we need slightly stronger condition:

Denote

$\|A_n\| = (\sum_{t,k=1}^n a_n(t,s)^2)^{1/2}$ Euclidean norm

$\|A_n\|_{sp} = \max_{\|x\|=1} \|A_n x\|$ spectral norm

CLT in zero diagonal case: $a_n(t,t) = 0$

Sufficient condition:

$$\frac{\|A_n\|_{sp}}{\|A_n\|} \rightarrow 0, \quad n \rightarrow \infty.$$

Then

$$\frac{T_n - E[T_n]}{\sqrt{\text{Var}(T_n)}} \rightarrow N(0, 1).$$

Discussed by: Rotar (1973), Jong (1987), Guttorp and Lockhart (1988), Mikosch (1991)

Note: 1. Condition implies CLT if $\|T_n\| \leq C$,

$$\text{Var}(T_n) = 2\|T_n\|^2 \rightarrow \infty$$

2. Only second finite moment is needed: $E\varepsilon_t^2 < \infty$.

CLT in non-zero diagonal case: $a_n(t, t) \neq 0$. More subtle, we discuss it later

Asymptotic theory for quadratic forms

Assume now that X_t are dependent linear variables

Objectives: 1. asymptotic normality theory for quadratic form $Q_{n,X}$ in dependent linear random variables X_t

$$Q_{n,X} = \sum_{k,t=1}^n d_n(k-t) X_k X_t$$

2. for the use in kernel, and other estimators converging at a rate not necessarily

$n^{1/2}$.

3. suitable for all types of dependence (long, short and negative memory) of X_t

4. conditions should be easy to verify

Known results for dependent X_k

1. The case $d_n(t) \equiv d(t)$. Conditions for CLT with normalization \sqrt{n} were derived in

Fox and Taqqu (1987), Avram (1988), Giraitis and Surgailis (1990) and others

4 finite moments needed, $EX_t^4 < \infty$, $E\varepsilon_t^4 < \infty$

Note Direct verification of CLT when $d_n(t)$ depends on n is difficult.

We wish to allow slow growth of $Var(Q_{n,X}) = o(n)$, to cover kernel estimation.

Method: we approximate $Q_{n,X}$ by a quadratic form

$$Q_{n,\varepsilon} = \sum_{k,t=1}^n e_n(k-t)\varepsilon_k\varepsilon_t,$$

in i.i.d. variables ε_t (innovations of $\{X_t\}$)

Note: The existing research, based on this method,

Phillips and Solo (1992), Mikosch (1995), Kokoszka and Taqqu (1996) deals with the case

$$d_n(t) \equiv d(t), \text{ four moments}$$

It provides only the bound

$$\text{Var}(Q_{n,X} - Q_{n,\varepsilon}) = o(n)$$

and CLT with normalization \sqrt{n} . Not good enough, to cover the case $Var(Q_{n,X}) = o(n)$.

Our approach:

1. Write

$$Q_{n,X} = Q_{n,\varepsilon} + [Q_{n,X} - Q_{n,\varepsilon}]$$

2. Use CLT for $Q_{n,\varepsilon}$ with non-vanishing diagonal

3. Main technical problem: show that

$$Q_{n,\varepsilon} \text{ dominates } [Q_{n,X} - Q_{n,\varepsilon}]$$

we need very sharp upper bound for

$$[Q_{n,X} - Q_{n,\varepsilon}]$$

Assumptions on X_t (linear process)

$$X_t = \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-j}.$$

Property: There exists $d \in (-1/2, 1/2)$ such that ψ_j satisfy

$$\psi_j = O(j^{-1+d}), \quad |\psi_j - \psi_{j+1}| = O(j^{-2+d}), \quad \text{if } d \neq 0$$

and, $\sum_{j=0}^{\infty} \psi_j = 0$, if $d < 0$.

If $d = 0$, then there exists $\alpha > 1$ such that

$$\sum_{j=n}^{\infty} |\psi_j| = O(n^{-\alpha}).$$

Example: the above property holds if $\{X_t\}$ is defined by

$$(1 - L)^d A(L)X_t = \varepsilon_t, \quad A(L) = \sum_{j=0}^{\infty} a_j L^j$$

where $-1/2 < d < 1/2$ is memory parameter, L is the lag operator, and AR coefficients decay fast:

$$a_j = O(r^j), \quad \text{for some } 0 < r < 1.$$

For example, ARFIMA(p,d,q) models. Then $\{X_t\}$ has spectral density

$$f(\lambda) = (2\pi)^{-1} \left| \sum_{j=0}^{\infty} \psi_j e^{-i\lambda j} \right|^2$$

with property

$$f(\lambda) = |\lambda|^{-2d} (b_0 + O(|\lambda|^2)), \quad \lambda \rightarrow 0.$$

Question: how to construct the approximating quadratic form $Q_{n,Z\varepsilon}$?

Denote by

$$I_n(\lambda) = \frac{1}{2\pi n} \left| \sum_{j=1}^n X_j e^{i\lambda j} \right|^2, \quad I_{n,\varepsilon}(\lambda) = \frac{1}{2\pi n} \left| \sum_{j=1}^n \varepsilon_j e^{i\lambda j} \right|^2$$

periodograms of $\{X_t\}$ and $\{\varepsilon_t\}$.

Under

Assumption on $d_n(t)$: there exist a real even function $\eta_n(\lambda)$ such that

$$d_n(t) = \int_{-\pi}^{\pi} \eta_n(\lambda) e^{i\lambda t} d\lambda$$

we can write

$$Q_{n,X} = \sum_{k,t=1}^n d_n(k-t)X_kX_t \equiv (2\pi n) \int_{-\pi}^{\pi} \eta_n(\lambda)I_n(\lambda)d\lambda$$

Bartlett decomposition

$$I_n(\lambda) = 2\pi f(\lambda)I_{n,\varepsilon}(\lambda) + \text{"small term"}$$

suggests that we can write

$$Q_{n,X} = Q_{n,\varepsilon} + \text{"small term"}$$

where

$$Q_{n,\varepsilon} = \sum_{k,t=1}^n e_n(k-t)\varepsilon_k\varepsilon_t \equiv (2\pi n) \int_{-\pi}^{\pi} \eta_n(\lambda)2\pi f(\lambda)I_{n,\varepsilon}(\lambda)d\lambda$$

here

$$e_n(t) = \int_{-\pi}^{\pi} (2\pi\eta_n(\lambda)f(\lambda))e^{i\lambda t}d\lambda.$$

Approximation of $Q_{n,X}$:

Objective: to derive sharp upper bounds of

$$\text{Var}(Q_{n,X} - Q_{n,\varepsilon}) \text{ and } E|Q_{n,X} - Q_{n,\varepsilon}|$$

Assumption on $\eta_n(\lambda)$: There exists $-1 < \beta < 1$ and $k_n \geq 0$,

$$|\eta_n(\lambda)| \leq k_n |\lambda|^{-\beta}, \quad \lambda \in [-\pi, \pi], \quad n \geq 1.$$

Note that the weight function in approximating form $Q_{n,\varepsilon}$, has the bound

$$|\eta_n(\lambda)| f(\lambda) \leq C k_n |\lambda|^{-(2d+\beta)}, \quad \lambda \rightarrow 0$$

Main approximation result

Theorem 2.1 Assume that

$$\delta := 2d + \beta < 1/2.$$

(i) If $E\varepsilon_t^4 < \infty$, then

$$[\text{Var}(Q_{n,X} - Q_{n,\varepsilon})]^{1/2} = O(r_n)$$

where

$$r_n = \begin{cases} k_n, & \text{if } d = 0, \\ k_n n^{\max(\delta, 0)} \log n, & \text{if } d \neq 0. \end{cases}$$

(ii) If $E\varepsilon_t^2 < \infty$, then

$$E|Q_{n,X} - Q_{n,\varepsilon}| = O(\bar{r}_n)$$

where

$$\bar{r}_n = \begin{cases} k_n, & \text{if } d = 0, \\ k_n n^{\max(\delta, d, 0)} \log n, & \text{if } d \neq 0 \end{cases}$$

(Note that $r_n \leq \bar{r}_n$.)

Comment: 1. Approximation rate depends on $\delta = 2d + \beta$. It allows memory compensation (d positive, β negative)

2. k_n plays a secondary role

3. Approximation precision is very high when δ is small or negative

4. In some case existing results are improved by $n^{-1/2}$. For example, if $d = 0$, $\eta_n(\lambda) \equiv \eta(\lambda)$ and $\beta = 0$, then Brockwell and Davis (1991) approximation is

$$[\text{Var}(Q_{n,X} - Q_{n,\varepsilon})]^{1/2} = o(n^{1/2})$$

Derived approximation is

$$[\text{Var}(Q_{n,X} - Q_{n,\varepsilon})]^{1/2} = O(1)$$

Central Limit theorem for Quadratic forms

To derive the CLT for $Q_{n,X}$ we have to

1. assume that the main term $Q_{n,\varepsilon}$ dominates the approximation error.

2, Use CLT for $Q_{n,\varepsilon}$

Notations: denote E_n the matrix $E_n = (e_n(t - k))_{t,k=1,\dots,n}$

$\|E_n\| = (\sum_{t,k=1}^n e_n^2(t - k))^{1/2}$ Euclidean norm.

Recall

$$|\eta_n(\lambda)| \leq k_n |\lambda|^{-\beta}$$

Theorem 2.2 Assume that $\beta + 2d < 1/2$ and $E\varepsilon_t^4 < \infty$. Suppose that

$$\frac{k_n}{\|E_n\|} \rightarrow 0, \quad \text{when } d = 0$$

$$\frac{k_n n^{\max(\beta+2d,0)} \log n}{\|E_n\|} \rightarrow 0, \quad \text{when } d \neq 0$$

then,

$$\text{Var}(Q_{n,X})/\text{Var}(Q_{n,\varepsilon}) \rightarrow 1, \quad \text{Var}(Q_{n,X}) \asymp \|E_n\|^2$$

and

$$(\text{Var}(Q_{n,X}))^{-1/2}(Q_{n,X} - EQ_{n,X}) \xrightarrow{d} N(0, 1).$$

Note: conditions of the CLT are comparable to those of the classical CLT in case of i.i.d. variables

Statistical applications involve the integrated periodograms

$$T_{n,X} = \int_{-\pi}^{\pi} \eta_n(\lambda) I_n(\lambda) d\lambda$$

Theorem 2.3 gives conditions when centering $ET_{n,X}$ can be replaced by explicit constant

$$ET_{n,\varepsilon} = \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda.$$

Recall that $\delta = \beta + 2d$,

$$\bar{r}_n = \begin{cases} k_n, & \text{if } d = 0, \\ k_n n^{\max(\delta, d, 0)} \log n, & \text{if } d \neq 0 \end{cases}$$

Theorem 2.3. Assume that $\delta < 1/2$ and

$$\frac{\bar{r}_n}{\|E_n\|} \rightarrow 0$$

(i) If $E\varepsilon_t^4 < \infty$, then

$$[\text{Var}(T_{n,X})]^{-1/2}(T_{n,X} - \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda) \xrightarrow{d} N(0, 1).$$

(ii) If

$$E\varepsilon_t^{2+\delta} < \infty \text{ and } \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda = o(n^{-1/2} \|E_n\|)$$

then

$$\text{Var}(T_{n,X}) = \frac{\|E_n\|^2}{2(\pi n)^2} (1 + o(1))$$

and

$$\frac{\sqrt{2}\pi n}{\|E_n\|} (T_{n,X} - \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda) \xrightarrow{d} N(0, 1).$$

(iii) If

$$E\varepsilon_t^2 < \infty \text{ and } \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda \equiv 0$$

then

$$\frac{\sqrt{2\pi n}}{\|E_n\|} (T_{n,X} - \int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda) \xrightarrow{d} N(0, 1).$$

Discussion of the results

Assumption on weights ψ_j can be replaced by a stronger condition

Assumption. There exists $d \in (-1/2, 1/2)$ and a constant $c \neq 0$ such that

$$\psi_j = \begin{cases} cj^{-1+d}(1 + O(j^{-1})), & \text{if } d \in (0, 1/2), \\ cj^{-1+d}(1 + O(j^{-1})) \text{ and } \sum_{j=0}^{\infty} \psi_j = 0, & \text{if } d \in (-1/2, 0). \end{cases}$$

Comment: 1. Assumption motivated by common time series models

2. It implies that the spectral density

$$f(\lambda) = c|\lambda|^{-2d}(1 + o(1)), \quad \text{as } \lambda \rightarrow 0,$$

3. Assumption $|\eta_n(\lambda)| \leq k_n|\lambda|^{-\beta}$ is weak, easy to check

4. In case $k_n \equiv K$, $2d + \beta \leq 0$ (for example, $d = 0$, $\beta = 0$), the bound

$$E|Q_{n,X} - Q_{n,\varepsilon}| = O(\log n),$$

is a much sharper bound than

$$E|Q_{n,X} - Q_{n,\varepsilon}| = O(n^{1/2}),$$

in Brockwell and Davis (1991), Kokoszka and Taqqu (1996,97)

5. In case $\eta_n(\lambda) \equiv \eta(\lambda)$, Taqqu, Fox (1986), Giraitis and Surgailis (1991) used the same condition

$$2d + \beta < 1/2$$

If $2d + \beta > 1/2$ CLT might not hold.

6. Applications show that CLT in kernel estimation, instead of 4, requires only $2+\delta$ moments

$$E\varepsilon_t^{2+\delta} < \infty$$

7. In applications, the main term $Q_{n,\varepsilon}$ dominates the remainder $Q_{n,X} - Q_{n,\varepsilon}$, and allows

$$\text{Var}(Q_{n,X}) = o(n).$$

CLT for quadratic forms of i.i.d. random variables

Problem: asymptotic normality result is based on approximation result, and normality of $Q_{n,\varepsilon}$.

Consider the general quadratic form

$$T_n = \sum_{t,k=1}^n a_n(t,k) \varepsilon_t \varepsilon_k$$

where $A_n = (a_n(t,k))_{t,k=1,\dots,n}$ a real symmetric matrix.

Comment: The case of zero diagonal $a_n(t,t) = 0$ is well investigated. Next theorem allows non-zero diagonal

Theorem 4.1 Assume that

$$\frac{\|A_n\|_{sp}}{\|A_n\|} \rightarrow 0.$$

(i) [**Non-zero diagonal**]. If $E\varepsilon_t^4 < \infty$, then

$$(\text{Var}(T_n))^{-1/2}(T_n - ET_n) \xrightarrow{d} N(0, 1).$$

(ii) [**Vanishing Diagonal**] If

$$E\varepsilon_t^{2+\delta} < \infty \quad (\text{for some } \delta > 0) \quad \text{and} \quad \sum_{t=1}^n a_n^2(t, t) = o(\|A_n\|^2),$$

then

$$\frac{1}{\sqrt{2}\|A_n\|}(T_n - ET_n) \xrightarrow{d} N(0, 1).$$

(iii) [**Zero diagonal**] If

$$E\varepsilon_t^2 < \infty \quad \text{and} \quad a_n(t, t) = 0, \quad t = 1, \dots, n,$$

then CLT(ii) is valid

Special case: A_n is a Toeplitz matrix with entries

$$a_n(t, k) = \int_{-\pi}^{\pi} e^{i(t-k)x} g_n(x) dx, \quad t, k = 1, \dots, n,$$

where $g_n(x)$, $|x| \leq \pi$ is an even real function.

Theorem 4.2 Let A_n be a Toeplitz matrix and for some $0 \leq \alpha < 1$

$$|g_n(\lambda)| \leq k_n |\lambda|^{-\alpha}, \quad n \geq 1.$$

(i) Then

$$\|A_n\|_{sp} \leq C k_n n^\alpha \quad n \geq 1.$$

(ii) If

$$\frac{k_n n^\alpha}{\|A_n\|} \rightarrow 0$$

then

$$\frac{\|A_n\|_{sp}}{\|A_n\|} \rightarrow 0.$$

Comment: 1. If $|g_n(\lambda)| \leq C$, and $Var(T_n) \rightarrow \infty$, then T_n satisfies CLT.

2. Condition on g is precise. For example, if

$$g_n(x) = |x|^{-\alpha}, \quad 0 \leq \alpha < 1$$

then

$$\|A_n\| \sim n^{\max(1/2, \alpha)}.$$

Hence, a) for $0 \leq \alpha < 1/2$,

$$\frac{\|A_n\|_{sp}}{\|A_n\|} \leq Cn^{\alpha-1/2} \rightarrow 0$$

and CLT holds.

b) if $1/2 < \alpha < 1$, then non-CLT holds (Giraitis, Taqqu, Terrin (1998))

Applications

CLT for quadratic forms is one of the main tools in inference of time series

A number of estimators/tests can be written as a quadratic form $Q_{n,X}$ (integrated periodogram with kernel $\eta_n(\lambda)$)

Important applications:

1. spectral estimation
2. kernel estimation
3. Whittle estimation
4. goodness-of-fit test

Example of application in kernel estimation

Illustration. Assume that $\{X_t\}$ is a linear short memory sequence with $d = 0$.

We estimate $f(0)$ using kernel estimator

$$\hat{f}(0) = \int_{-\pi}^{\pi} \eta_n(\lambda) I_n(\lambda) d\lambda$$

where

$$\eta_n(\lambda) = (2\pi q)^{-1} \left| \sum_{j=1}^q e^{ij\lambda} \right|^2$$

is the Fejér kernel.

The estimator $\hat{f}(0)$ uses the Bartlett window. The bandwidth

$$q \rightarrow \infty, \quad q = o(n), \quad \text{as } n \rightarrow \infty.$$

Existing results: Anderson (1994), Theorem 9.4.1, shows that

$$(n/q)^{1/2}(\hat{f}(0) - E\hat{f}(0)) \rightarrow N(0, V^2)$$

It assumes finite fourth moment

$$E\varepsilon_t^4 < \infty.$$

Note: 1. Centering by $E[\hat{f}(0)]$ is not convenient,

2. analysis of the bias $f(0) - E[\hat{f}(0)]$ difficult.

Our results: 1. imply asymptotic normality of $\hat{f}(0)$:

a) under $2 + \delta$ moments, $E\varepsilon_t^{2+\delta} < \infty$.

b) allows simple deterministic centering

Assumptions: Assume that f is continuous and

$$f(\lambda) = f(0) + O(\lambda^2), \text{ as } \lambda \rightarrow 0.$$

Since

$$|\eta_n(\lambda)| \leq Cq$$

then

$$|\eta_n(\lambda)| \leq k_n |\lambda|^{-\beta}$$

So,

$$k_n = Cq, \quad \beta = 0, \quad \text{and} \quad \delta = \beta + 2d = 0.$$

It is straightforward to check that

$$\|E_n\|^2 = \sum_{t,k=1}^n e_n^2(t-k) = \int_{-\pi}^{\pi} \left| \sum_{j=1}^n e^{it(x+y)} \right|^2 \eta_n(x) f(x) \eta_n(y) f(y) dx dy$$

$$\sim qn (8/3)\pi^2 f(0)^2,$$

Then

$$\frac{\bar{r}_n}{\|E_n\|} = \frac{k_n}{\|E_n\|} \sim C \frac{q}{\sqrt{qn}} = C \frac{\sqrt{q}}{\sqrt{n}} \rightarrow 0.$$

Since f is continuous, then

$$\int_{-\pi}^{\pi} \eta_n(\lambda) f(\lambda) d\lambda = (2\pi q)^{-1} \int_{-\pi}^{\pi} \left| \sum_{j=1}^q e^{ij\lambda} \right|^2 f(\lambda) d\lambda \rightarrow f(0) = o(n^{-1/2} \|E_n\|)$$

since $n^{-1/2} \|E_n\| \sim cq^{1/2} \rightarrow \infty$. Hence

$$(n/q)^{1/2} (\hat{f}(0) - \int_{-\pi}^{\pi} \eta_n(x) f(x) dx) \xrightarrow{d} N(0, \frac{4}{3} f^2(0)).$$

Note: this convergence does not follow from any existing CLT's for quadratic forms of linear processes because

a) it involves rate of convergence different than \sqrt{n}

b) function η_n depends on n .

c) condition $f(\lambda) = f(0) + O(\lambda^2)$ allows to obtain the upper bound of the bias:

$$\int_{-\pi}^{\pi} \eta_n(x) f(x) dx - f(0) = O(q^{-1}).$$