

# Infinite variance stable limits for dependent sequences <sup>1</sup>

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## STABLE LIMITS FOR AN IID SEQUENCE

- For an iid real-valued sequence  $(X_t)$  consider the partial sums
 
$$S_n = X_1 + \cdots + X_n, n \geq 1.$$
- Using classical limit theory for sums of independent random variables, e.g. Gnedenko, Kolmogorov (1954), Feller (1971), Petrov (1975, 1996), one can show that there exist sequences  $0 < a_n \rightarrow \infty$  and  $b_n \in \mathbb{R}$  and a random variable  $Y$  with non-degenerate law  $H$  such that

$$a_n^{-1}(S_n - b_n) \xrightarrow{d} Y \sim H$$

if and only if **either**  $f(x) = EX^2 I_{\{|X| \leq x\}}$ ,  $x > 0$ , is slowly varying **or**  $X$  is regularly varying with index  $\alpha \in (0, 2)$ , i.e., there exist  $p, q \geq 0$  with  $p + q = 1$  and a slowly varying

function  $L$  such that

$$P(X > x) \sim p \frac{L(x)}{x^\alpha} \quad \text{and} \quad P(X \leq -x) \sim q \frac{L(x)}{x^\alpha},$$

- $H = H_\alpha$ ,  $\alpha \in (0, 2]$ , is  $\alpha$ -stable in the convolution sense, i.e. for any  $n \geq 2$  and an iid sequence  $(Y_t)$  with common distribution  $H$ , there exist  $c_n > 0$  and  $d_n \in \mathbb{R}$  such that

$$c_n^{-1}(Y_1 + \cdots + Y_n - d_n) \stackrel{d}{=} Y.$$

- Moreover, for  $\alpha \in (0, 2)$ ,  $(a_n)$  can be chosen such that

$$P(|X| > a_n) \sim n^{-1} \quad \text{and} \quad b_n = n EXI_{\{|X| \leq a_n\}}.$$

- Classical proofs are based on characteristic function arguments.
- An alternative way of proving this result goes back to [LePage](#),

[Woodroffe](#), [Zinn \(1981\)](#), [Resnick \(1986\)](#); see also [Resnick \(2007\)](#).

- Since regular variation of  $X$  for any  $\alpha > 0$  is equivalent to the weak convergence of the point processes

$$N_n = \sum_{t=1}^n \varepsilon_{a_n^{-1} X_t} \xrightarrow{d} N = \sum_{t=1}^{\infty} \varepsilon_{J_t} \sim \text{PRM}(\mu)$$

for some Poisson random measure  $N$  with mean measure  $\mu$  on  $\overline{\mathbb{R}} \setminus \{0\}$  given by

$$\mu(dx) = [p x^{-\alpha} I_{\{x>0\}} + q |x|^{-\alpha} I_{\{x<0\}}] dx .$$

- The mapping  $T_\epsilon : M_p \rightarrow \mathbb{R}$  given by

$$T_\epsilon(m) = T_\epsilon\left(\sum_t \varepsilon_{j_t}\right) = \sum_t j_t I_{\{|j_t|>\epsilon\}}$$

is a.s. continuous relative to the distribution of  $N$  for every  $\epsilon > 0$ .

- Hence

$$T_\epsilon(N_n) = \sum_{t=1}^n (a_n^{-1} X_t) I_{\{|a_n^{-1} X_t| > \epsilon\}} \xrightarrow{d} T_\epsilon(N) = \sum_{t=1}^{\infty} J_t I_{\{|J_t| > \epsilon\}}.$$

- For  $\alpha \in (0, 2)$  the right-hand side has a limit as  $\epsilon \downarrow 0$  (with additional centering for  $\alpha \in [1, 2)$ ): *series representation of an  $\alpha$ -stable random variable.*

- **Example.** Assume  $p = 1$  ( $X$  is totally skewed to the right) and  $\alpha \in (0, 1)$ . Then  $N = \sum_{t=1}^{\infty} \varepsilon_{\Gamma_t^{-1/\alpha}}$ , where  $0 < \Gamma_1 < \Gamma_2 < \dots$  are the points of a homogeneous Poisson process. Hence

$$T_\epsilon(N) = \sum_{t=1}^{\infty} \Gamma_t^{-1/\alpha} I_{\{|\Gamma_t^{-1/\alpha}| > \epsilon\}} \xrightarrow{\text{a.s.}} \sum_{t=1}^{\infty} \Gamma_t^{-1/\alpha} \quad \text{as } \epsilon \downarrow 0.$$

which represents an  $\alpha$ -stable random variable.

- It finally suffices to show that

$$\lim_{\epsilon \downarrow 0} \limsup_{n \rightarrow \infty} P(|a_n^{-1} S_n - T_\epsilon(N_n) - E(\cdot)| > \delta) = 0, \quad \delta > 0,$$

e.g. by showing that  $\text{var}(a_n^{-1} S_n - T_\epsilon(N_n))$  can be made small.

## GENERALIZATIONS TO DEPENDENT SEQUENCES

### Linear processes.

- Recall the definition of a **linear process**

$$(1) \quad X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j}, \quad t \in \mathbb{Z},$$

for sequences of suitable constants  $\psi_j$ ,  $j \in \mathbb{Z}$ , and an iid sequence  $(Z_t)$ .

- If  $Z$  is regularly varying with index  $\alpha > 0$ , i.e.,

$$P(Z > x) \sim p \frac{L(x)}{x^\alpha} \quad \text{and} \quad P(Z \leq -x) \sim q \frac{L(x)}{x^\alpha},$$

and the series (1) converges a.s. then  $X$  is regularly varying with index  $\alpha > 0$ .<sup>2</sup>

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<sup>2</sup>The converse is not true in general; see Jacobsen, Mikosch, Rosiński, Samorodnitsky (2009).

- In a series of papers, [Davis and Resnick \(1985, 1986\)](#) proved that the sequence of the partial sums  $(a_n^{-1}S_n)$  has a stable limit for  $\alpha \in (0, 2)$ . They also showed the joint convergence for

$$\sum_{t=1}^n \left( a_n^{-1}X_t, a_n^{-1}X_t^2, \tilde{a}_n^{-1}X_tX_{t+1}, \dots, \tilde{a}_n^{-1}X_tX_{t+h} \right) - b_n$$

towards a mixed stable distribution. This was achieved by using the weak convergence of the underlying point processes and a continuous mapping argument.

- [Phillips and Solo \(1992\)](#) used the structure of a linear process to show that, under general weak dependence conditions,

$$a_n^{-1} \left( \sum_{t=1}^n X_t - \sum_{j=0}^n \psi_j \sum_{t=1}^n Z_t \right) \xrightarrow{P} 0,$$

thus the stable CLT for  $(X_t)$  follows from the one for  $(Z_t)$ .



- [Kasahara, Maejima, Vervaat \(1988\)](#) also considered stable FCLTs in the case of strong dependence.

## Mixing conditions.

- Let  $(X_t)$  be a strictly stationary sequence with partial sum process  $S_n = X_1 + \cdots + X_n$ ,  $n \geq 1$ .
- Davis and Hsing (1995) proved stable limit theory by using the point process approach.
- Davis and Hsing (1995) require the mixing condition  $\mathcal{A}(a_n)$  in terms of the point processes

$$N_{nm} = \sum_{t=1}^m \varepsilon_{a_n^{-1} X_t} \quad \text{and} \quad N_n = N_{nn} = \sum_{t=1}^n \varepsilon_{a_n^{-1} X_t}.$$

- They require closeness of the Laplace functionals

$$E e^{-\int f dN_n} - \left( E e^{-\int f dN_{nm}} \right)^{k_n} \rightarrow 0,$$

where  $m = m_n \rightarrow \infty$ ,  $k_n = [n/m] \rightarrow \infty$ .

- [Bartkiewicz et al. \(2010\)](#) prove stable limit theory by using characteristic functions.
- [Bartkiewicz et al. \(2010\)](#) require a mixing condition in terms of the characteristic functions

$$\varphi_n(x) = Ee^{ixa_n^{-1}S_n} \quad \text{and} \quad \varphi_{nm}(x) = Ee^{ixa_n^{-1}S_m} .$$

- They require closeness of the characteristic functions

$$\varphi_n(x) - \left(\varphi_{nm}(x)\right)^{k_n} \rightarrow 0 ,$$

where  $m = m_n \rightarrow \infty$ ,  $k_n = [n/m] \rightarrow \infty$ .

- Conditions of this type as well as  $\mathcal{A}(a_n)$  follow from strong mixing with suitable rates.

- These conditions imply that the corresponding limits, if they exist, are *infinitely divisible*.

## Conditions on the tails.

- To ensure convergence to an infinite variance stable limit, we require **regular variation** of the finite-dimensional distributions of  $(X_t)$  as in Davis and Hsing (1995) and Bartkiewicz et al. (2010):<sup>3</sup> There exist  $\alpha \geq 0$  and, for every  $h \geq 1$ , a non-constant vector  $\Theta_h$  on the unit sphere of  $\mathbb{R}^h$  such that for  $Y_h = (X_1, \dots, X_h)$ , as  $x \rightarrow \infty$ ,

$$\frac{P(|Y_h| > xc)}{P(|Y_h| > x)} \rightarrow c^{-\alpha}, \quad c > 0,$$

and

$$P(Y_h/|Y_h| \in \cdot \mid |Y_h| > x) \xrightarrow{w} P(\Theta_h \in \cdot).$$

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<sup>3</sup>Regular variation is not necessary for partial sum convergence of a strictly stationary sequence; Surgailis (2004), Gouëzel (2004)

- We say that  $(X_t)$  is **regularly varying with index  $\alpha > 0$** .
- An *equivalent definition* is the following: for every  $h \geq 1$ , there exists a non-null Radon measure  $\mu_h$  on  $\overline{\mathbb{R}}^h \setminus \{0\}$  such that

$$n P(a_n^{-1} Y_h \in \cdot) \xrightarrow{v} \mu_h(\cdot),$$

where  $(a_n)$  satisfies  $P(|X| > a_n) \sim n^{-1}$ .

- The measure  $\mu_h$  satisfies  $\mu_h(t A) = t^{-\alpha} \mu_h(A)$ ,  $t > 0$ , for some  $\alpha \geq 0$ .

- **Examples.** Infinite variance stable stationary processes.

ARMA/linear processes with iid regularly varying noise.

Stochastic recurrence equations  $X_t = A_t X_{t-1} + B_t$  with iid non-negative  $((A_t, B_t))$  Kesten (1973), Goldie (1991).

GARCH processes  $X_t = \sigma_t Z_t$  with iid noise  $(Z_t)$  with infinite support.

Stochastic volatility processes with regularly varying noise  $(Z_t)$ .

Transformed Gaussian stationary sequence such that the one-dimensional marginals are regularly varying.

- If  $(X_t)$  is regularly varying with index  $\alpha > 0$  so are the linear combinations of any finite segment of this sequence: for  $A$  bounded away from zero with a smooth boundary,

$$n P(a_n^{-1} S_d \in A) \rightarrow \mu_d(\{x \in \mathbb{R}^d : x_1 + \cdots + x_d \in A\}).$$

- In particular, for  $d \geq 1$ ,

$$P(S_d > x) \sim p(d) P(|X| > x) \quad \text{and} \quad P(S_d \leq -x) \sim q(d) P(|X| > x).$$

- $(p(d))_{d \geq 1}$  and  $(q(d))_{d \geq 1}$  measure the strength of dependence in  $(X_t)$  with respect to the tails of partial sums.
- **Example.** For  $(X_t)$  iid and  $X > 0$ ,  $P(S_d > x) \sim d P(X > x)$ .

$$\text{For } X_t = X > 0, P(S_d > x) = P(dX > x) \sim d^\alpha P(X > x).$$



## MAIN RESULT

### Assumptions.

- The strictly stationary sequence  $(X_t)$  is mixing in the sense

$$\varphi_n(x) - \left(\varphi_{nm}(x)\right)^{k_n} \rightarrow 0, \quad x \in \mathbb{R},$$

where  $m = m_n \rightarrow \infty$ ,  $k_n = [n/m] \rightarrow \infty$ .

- $(X_t)$  is regularly varying with index  $\alpha \in (0, 2)$
- An anti-clustering and a centering condition hold.
- The following limits exist<sup>4</sup>

$$\text{(Jak)} \quad p = \lim_{d \rightarrow \infty} [p(d) - p(d-1)] \quad \text{and} \quad q = \lim_{d \rightarrow \infty} [q(d) - q(d-1)].$$

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<sup>4</sup>This condition was introduced in Jakubowski (1993,1997).

- Then  $p, q \geq 0$  and for  $(a_n)$  with  $P(|X| > a_n) \sim n^{-1}$ ,  
 $a_n^{-1} S_n \xrightarrow{d} Y_\alpha$ , where  $Y_\alpha$  is  $\alpha$ -stable with characteristic function  
 $\psi_\alpha$ <sup>5</sup> given by
 
$$-\log \psi_\alpha(x)$$

$$= |x|^\alpha \frac{\Gamma(2 - \alpha)}{1 - \alpha} ((p + q) \cos(\pi\alpha/2) - i \operatorname{sign}(x) (p - q) \sin(\pi\alpha/2))$$

$$= \chi_\alpha(x, p, q), \quad x \in \mathbb{R}.$$

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<sup>5</sup>Shown for  $\alpha \neq 1$  only.

- The condition (Jak) implies that  $p = \lim_{d \rightarrow \infty} d^{-1}p(d)$  and  $q = \lim_{d \rightarrow \infty} d^{-1}q(d)$  exist.
- **Examples.  $m_0$ -dependence:**  $p = p(m_0 + 1) - p(m_0)$ ,  
 $q = q(m_0 + 1) - q(m_0)$ .

**Stochastic volatility model:**  $X_t = \sigma_t Z_t$  with stationary Gaussian  $\log \sigma_t$  and iid regularly varying  $(Z_t)$ :  $p = dp - (d - 1)p$  and  $q = dq - (d - 1)q$ .

**Stochastic recurrence equations:**  $X_t = A_t X_{t-1} + B_t$  with iid non-negative  $((A_t, B_t))$ . Let  $E[A^\kappa] = 1$  have the (unique) solution  $\alpha > 0$ . Then  $(X_t)$  is regularly varying with index  $\alpha$

and

$$P(X > x) \sim c_0 x^{-\alpha}, \quad x \rightarrow \infty.$$

- With  $\Pi_t = A_1 \cdots A_t$ ,  $t \geq 1$ ,

$$(X_1, \dots, X_d) = X_0 (\Pi_1, \dots, \Pi_d) + R_d$$

where  $X_0$  is independent of  $R_d, \Pi_1, \dots, \Pi_d$ .

- Hence, with  $T_d = \sum_{i=1}^d \Pi_i$ , by a result of Breiman (1965)

$$P(S_d > x) \sim P(X_0 T_d > x) \sim P(X_0 > x) E[T_d^\alpha]$$

and  $p(d) = E[T_d^\alpha]$ . Since  $E[A^\alpha] = 1$ ,

$$\begin{aligned} p(d+1) - p(d) &= E[T_{d+1}^\alpha] - E[T_d^\alpha] = E[A_{d+1}^\alpha (1 + T_d)^\alpha] - E[T_d^\alpha] \\ &= E[(1 + T_d)^\alpha - T_d^\alpha] \rightarrow E[(1 + T_\infty)^\alpha - T_\infty^\alpha]. \end{aligned}$$

- Although  $E[T_\infty^\alpha] = \infty$ ,

$$d^{-1}E[T_d^\alpha] = E[d^{-1/\alpha}T_d]^\alpha \rightarrow E[(1 + T_\infty)^\alpha - T_\infty^\alpha] < \infty.$$

- **Squared GARCH processes** can be embedded in stochastic recurrence equations. Similar results hold for  $(X_t^2)$  and  $(\sigma_t^2)$  and also for  $(X_t)$ .

## MAIN IDEA OF PROOF

- In view of the mixing condition it follows that  $(a_n^{-1}S_n)$  has the same limit as  $(a_n^{-1} \sum_{i=1}^m S_{mi})$ , where  $S_{mi}$ ,  $i = 1, \dots, k_n$ , are iid copies of  $S_m$ .

- For this triangular array, it suffices to show that

$$k_n (\varphi_{nm}(x) - 1) = k_n \log \varphi_{nm}(x) + o(1) \rightarrow \log \psi_\alpha(x) = -\chi_\alpha(x, p, q).$$

- **Key lemma.** Under regular variation of  $(X_t)$  and with the anti-clustering condition,

$$\lim_{d \rightarrow \infty} \limsup_{n \rightarrow \infty} \left| k_n (\varphi_{nm}(x) - 1) - n (\varphi_{nd}(x) - \varphi_{n,d-1}(x)) \right| = 0, \quad x \in \mathbb{R}.$$

- Under regular variation of  $S_d$ ,

$$n (\varphi_{nd}(x) - 1) \rightarrow -\chi_\alpha(x, p(d), q(d)), \quad x \in \mathbb{R},$$

and

$$\begin{aligned} & \chi_\alpha(x, p(d), q(d)) - \chi_\alpha(x, p(d-1), q(d-1)) \\ &= \chi_\alpha(x, p(d) - p(d-1), q(d) - q(d-1)) \\ &\rightarrow \chi_\alpha(x, p, q) \end{aligned}$$

## RELATED WORK

- [Balan and Louhichi \(2009\)](#) use the point process process approach for partial sums of triangular arrays of dependent random variables to show convergence towards infinitely divisible laws.
- [Buraczewski, Damek, Guivarc'h \(2009,2010\)](#) prove limit theory for multivariate stochastic recurrence equations  $X_t = A_t X_{t-1} + B_t$  without extra mixing conditions.
- [Tyran-Kamińska \(2010\)](#) proves a FCLT with stable Lévy motion under the condition

$$P(|X_j| > x \mid |X_0| > x) \rightarrow 0, \quad j \geq 1,$$

which is necessary under the  $J_1$ -topology.



- [Basrak, Krizmanić and Segers \(2010\)](#) prove a FCLT with stable Lévy motion in the  $M_1$ -topology under  $\mathcal{A}(a_n)$  and using the point process approach.



FIGURE 1. Til lykke.